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On the center of mass of Ising vectors

M. COPELLI, M. BOUTEN, C. VAN DEN BROECK and B. VAN ROMPAEY Limburgs Universitair Centrum - B-3590 Diepenbeek, Belgium

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Abstract. – We show that the center of mass of Ising vectors that obey some simple constraints, is again an Ising vector.

Many problems in statistical mechanics are formulated in terms of an N-dimensional vector \mathbf{J} , with components J_i , i = 1, ..., N that take only binary values $J_i = \pm 1$. We will call such a vector an Ising vector. Its components represent for example a spin state (Ising model [1]), the occupancy of a site, etc. In the thermodynamic limit $N \to \infty$, only a subset of all the possible configurations $\{\mathbf{J}\}$ are typically realized. In many cases, they are characterized by simple constraints of the form

$$\lim_{N \to \infty} \frac{\mathbf{J} \cdot \mathbf{B}}{N} = R, \qquad \lim_{N \to \infty} \frac{\mathbf{J} \cdot \mathbf{J}'}{N} = q, \tag{1}$$

where \mathbf{J} and \mathbf{J}' are typical members of the subset, \mathbf{B} is a symmetry-breaking direction (imposed from the outside or arising through a phase transition), while q and R are physical properties describing the resulting macroscopic state (for example the magnetization or the density). In this letter, we focus on the center of mass of the vectors \mathbf{J} , that satisfy the above constraints. We report the surprising finding that it is an Ising vector whenever \mathbf{B} is Ising.

To construct the center of mass we follow a Monte Carlo approach by choosing at random n vectors \mathbf{J}^{a} , a = 1, ..., n, that satisfy the constraints, and considering their center of mass:

$$\mathbf{Y} = C^{-1} \sum_{a} \mathbf{J}^{a},\tag{2}$$

with the proportionality factor $C = \sqrt{n + n(n-1)q}$, so that the normalization condition $\mathbf{Y}^2 = N$ is obeyed. In general the vector \mathbf{Y} is not Ising, but our contention is that it becomes Ising in the limit $n \to \infty$, provided \mathbf{B} is Ising. In view of the permutation symmetry between the coordinate axes, it will be sufficient to prove that $B_1Y_1 = C^{-1}\sum_a B_1 J_1^a$ only takes the values +1 and -1 in this limit. To show that this is the case, we focus our attention on $\widehat{\mathbf{C}}$ EDP Sciences

the probability density P(y) of the variable $y = n^{-1} \sum_{a} B_1 J_1^a$, which differs by a factor $n^{-1}\sqrt{n+n(n-1)q} \xrightarrow{n \to \infty} \sqrt{q}$ from B_1Y_1 . It is given by

$$P(y) \sim \int \left[\prod_{a}^{n} \mathrm{d}\mathbf{J}^{a} P_{b}(\mathbf{J}^{a}) \,\delta(\mathbf{J}^{a} \cdot \mathbf{B} - NR)\right] \times \\ \times \left[\prod_{a < b} \delta(\mathbf{J}^{a} \cdot \mathbf{J}^{b} - Nq)\right] \delta\left(y - n^{-1}\sum_{a} B_{1} J_{1}^{a}\right) , \qquad (3)$$

where P_b is the measure restricting to vectors with binary components,

$$P_b(\mathbf{J}) = \prod_{j=1}^N \left[\frac{1}{2} \delta(J_j - 1) + \frac{1}{2} \delta(J_j + 1) \right] , \qquad (4)$$

and the proportionality constant has to be determined from the normalization condition $\int_{-\infty}^{\infty} dy P(y) = 1$. The r.h.s. of (3) resembles an ordinary replica calculation [2], but with as limit of interest the number of replicas *n* tending to infinity.

Rather than following the standard but lengthy calculations that are usual in this case, we present a more elegant, direct and expedient procedure. Since $y = n^{-1} \sum_{a} x_{a}$, with $x_{a} = J_{1}^{a}B_{1}$, we evaluate the joint probability density $P(\mathbf{x})$ of the *n*-dimensional vector with binary components x_{a} , $a = 1, \ldots, n$. Since all choices of the vectors \mathbf{J}^{a} that satisfy the constraints can be realized, the Shannon entropy is maximized under the constraints (1). Hence $P(\mathbf{x})$ is found by maximizing [3] its Shannon entropy $-\sum_{\mathbf{x}} P(\mathbf{x}) \ln P(\mathbf{x})$, subject to the constraints

$$\langle x_a \rangle = \frac{\mathbf{J}^a \cdot \mathbf{B}}{N} = R,$$

$$\langle x_a x_b \rangle = \frac{\mathbf{J}^a \cdot \mathbf{J}^b}{N} = q, \qquad (a < b).$$
 (5)

One finds

$$P(\mathbf{x}) = Z^{-1} \exp\left[\sum_{a} \hat{R}_a x_a + \sum_{a < b} \hat{q}_{ab} x_a x_b\right] , \qquad (6)$$

where Z follows from the normalization of $P(\mathbf{x})$. The values of the Lagrange multipliers $\{\hat{R}_a\}$ and $\{\hat{q}_{ab}\}$ have to be determined from the constraints (5). In view of the permutation symmetry in the replica indices, \hat{R}_a and \hat{q}_{ab} must be independent of a and b, $\hat{R}_a = \hat{R}$, $\hat{q}_{ab} = \hat{q}$, rendering the evaluation of Z very simple:

$$Z(\hat{R},\hat{q}) = e^{-n\hat{q}/2} \int Dz \left[\cosh\left(\hat{R} + z\sqrt{\hat{q}}\right) \right]^n , \qquad (7)$$

while (5), determining \hat{R} and \hat{q} , reduce to

$$R = \frac{1}{n} \frac{\partial}{\partial \hat{R}} \ln Z = \frac{\int \mathrm{d}u \, \exp\left[-(u-\hat{R})^2/2\hat{q}\right] (\cosh u)^n \, \tanh u}{\int \mathrm{d}u \, \exp\left[-(u-\hat{R})^2/2\hat{q}\right] (\cosh u)^n},$$
$$q = \frac{2}{n(n-1)} \frac{\partial}{\partial \hat{q}} \ln Z = \frac{\int \mathrm{d}u \, \exp\left[-(u-\hat{R})^2/2\hat{q}\right] (\cosh u)^n \, \tanh^2 u}{\int \mathrm{d}u \, \exp\left[-(u-\hat{R})^2/2\hat{q}\right] (\cosh u)^n}.$$
(8)



Fig. 1. – In order to account for a logarithmic divergence in the limit $q \to 1$, $\exp[-\gamma]$, $\gamma = n\hat{q}$, is plotted as a function of q for several values of n.

As a result of the "replica symmetry", we conclude from (6) that $P(\mathbf{x})$ is in fact a function of $\sum_a x_a = ny$. Hence P(y) is obtained from $P(\mathbf{x})$ by multiplication with a combinatorial factor, expressing the freedom to choose which n(1+y)/2 components of \mathbf{x} are +1 and which remaining ones are -1, for a given total value ny of their sum:

$$P(y) \sim \binom{n}{\frac{n(1+y)}{2}} \exp\left[n\hat{R}y + \frac{n^2\hat{q}y^2}{2}\right].$$
(9)

y can take the values $-1, -1 + 2/n, \ldots, 1 - 2/n, 1$, and the proportionality constant is again fixed by normalization. This result is in agreement with a direct evaluation of (3) but very different from the result for continuous components discussed in [4]. Unfortunately, the above expression is quite complicated, especially in view of the fact that we did not succeed in solving explicitly eqs. (8) determining the Lagrange multipliers. Concordantly, the components of **Y** are not binary for any finite n. As an illustration, we have included in fig. 1 the results obtained by a numerical solution of (8) for the special case q = R and several values of n. The corresponding results for the probability density for y (or equivalently, B_1Y_1), are plotted in fig. 2.

In order to extract the asymptotic behavior for the $n \to \infty$ limit, one needs to guess the asymptotic dependence on n of the Lagrange parameters. The correct scaling appears quite naturally in the calculations for the simpler case of vectors \mathbf{J} with continuous components. Here, we just note by inspection that eqs. (8) for the properly scaled Lagrange parameters $\rho \equiv n \hat{R}$ and $\gamma \equiv n \hat{q}$ read

$$R = \frac{\int \mathrm{d}u \, e^{-n\phi(u)} \sinh(u\rho/\gamma) \tanh u}{\int \mathrm{d}u \, e^{-n\phi(u)} \cosh(u\rho/\gamma)},$$
$$q = \frac{\int \mathrm{d}u \, e^{-n\phi(u)} \cosh(u\rho/\gamma) \tanh^2 u}{\int \mathrm{d}u \, e^{-n\phi(u)} \cosh(u\rho/\gamma)},$$
(10)

where

$$\phi(u) \equiv \frac{u^2}{2\gamma} - \ln \cosh u \;. \tag{11}$$



Fig. 2. – Probability density for B_1Y_1 according to eq. (9) for q = R and several values of n. The legend is the same for both plots.

The appearance of the hyperbolic functions of $u\rho/\gamma$ in eqs. (10) is due to the fact that ϕ is even. The saddle point approximation can now, for $n \to \infty$, be applied in a straightforward manner on the *u*-integrations, leading to the following simple and explicit solutions for the scaled Lagrange variables:

$$\gamma = \frac{\operatorname{arctanh}\sqrt{q}}{\sqrt{q}},$$

$$\rho = \frac{\operatorname{arctanh}(R/\sqrt{q})}{\sqrt{q}}.$$
(12)

Inserting this result together with the asymptotic expression for the combinatorial factor in (9), one finally obtains the following asymptotic result for P(y):

$$P(y) \sim \exp[\rho \ y] \exp n \left[\gamma \ y^2/2 - \ln \sqrt{1 - y^2} - y \operatorname{arctanh} y \right]$$

$$\stackrel{n \to \infty}{\to} \frac{1}{2} \left(1 + \frac{R}{\sqrt{q}} \right) \delta(y - \sqrt{q}) + \frac{1}{2} \left(1 - \frac{R}{\sqrt{q}} \right) \delta(y + \sqrt{q}) .$$
(13)

In view of the aforementioned relation between y and the (first) component of \mathbf{Y} , the convergence of the latter to an Ising vector follows immediately.

As a first application of the above result, we turn to the case of an Ising spin system in the ferromagnetic phase. Choosing the vector **B** with all its components equal to 1, we note that R plays the role of the magnetization. All the spin states **J** are otherwise allowed, and lie on the rim of the N-dimensional sphere with radius \sqrt{N} at fixed angle $\arccos R$ with **B**. It is thus clear that the center of mass is **B** itself. This trivial result is recovered from (13) by noting that $q = R^2$ in this case. Note that the constraint $\mathbf{J}^a \cdot \mathbf{J}^b = Nq$ is therefore redundant, which implies $\hat{q} = 0$ and $\hat{R} = \operatorname{arctanh} R$ (a result valid for any n).

A case of special symmetry is q = R. In this scenario, no macroscopic measure allows to distinguish between the symmetry-breaking direction **B** and each of the vectors \mathbf{J}^{a} . The Lagrange parameters also present the symmetry $\hat{R} = \hat{q}$, which can be seen from eqs. (8).

A third case of interest is the limit $R \to 0$, while q remains finite. In this case, the **J**-vectors lie in the subspace orthogonal to **B** and satisfy as single constraint the prescribed mutual overlap q. From eqs. (8) it is clear that $\hat{R} = 0$ is automatically satisfied, and one concludes



Fig. 3. – Probability density for B_1Y_1 in the $q = R^2$ scenario and several values of n. The lines represent the theoretical curve (9), while points represent results from a simulation of the mean-field ferromagnetic Ising model with N = 100, dimensionless temperature 1.09 and dimensionless magnetic field 0.1 (amounting to a magnetization $R \simeq 0.5$ —see text for details). Inset: second cumulant (σ^2) of the distribution as a function of 1/n. The squares represent the simulations, while the dashed straight line corresponds to the theory (in the asymptotic limit $n \to \infty$, eq. (13) implies $\sigma^2 \sim (1-R^2)/(nR^2)$).

from (13) that the center of mass of the **J**-vectors is again an Ising vector (but its components are equally likely to be +1 or -1).

It is very tempting to apply the results above to neural network learning problems [5] where a student perceptron \mathbf{J} learns from examples generated by a teacher perceptron \mathbf{B} [6-8]. Indeed, the so-called Gibbs learning [9] presents the symmetry q = R, and the interest of the center of mass is that it is, according to a simple general argument [10], see also [11], the "best" student having the largest overlap with the teacher (namely \sqrt{R}). Accordingly, R = 0 and $q \neq 0$ are constraints satisfied by Ising vectors which solve the capacity problem [12]. However, these are disordered systems, and the conditions on R and q alone do not convey all the information which is necessary to describe the constraints in \mathbf{J} space. Therefore, even though the constraints (1) are satisfied in neural network problems, result (13) does not apply to them.

We conclude with a verification of the theoretical prediction (9) by running simulations for the mean-field ferromagnetic Ising spin model, fig. 3. The Metropolis algorithm was allowed to run for a number of Monte Carlo steps per site (MCS) until thermalization was considered to be achieved. Then vectors were sampled every 5 MCS (to allow sufficient decorrelation between consecutive samplers) and summed to construct the center of mass. The small discrepancy for N = 100 with the theoretical prediction is due to finite-size effects. For N = 1000, the results are nearly indistinguishable on the scale of the figure from the theoretical values. It is interesting to note that the hard constraints of eq. (1) are satisfied *only* in the thermodynamic limit. In the simulations R (and q) are distributed with peaks whose width scales with $N^{-1/2}$. Nonetheless the effect of these fluctuations on the resulting $P(B_1Y_1)$ is negligible.

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